STUDIES



Surface curvature analysis of bivariate normal distribution: A Covid-19 data application on Turkey

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Keywords:

principal curvatures, Gaussian curvature, mean curvature, bivariate normal distribution, correlation Principal curvatures have free-form rigid surfaces' invariant features. Therefore they are widely used in several fields for various applications, such as determining the corresponding points between an object and a free-form scene. In this study, the authors analysed the surface curvature of a bivariate normal distribution. A novel approach for classifying bivariate normal surfaces based on curvature statistics concerning correlation structures is presented. The principal curvatures, Gaussian, and mean curvatures were obtained using the data generated from the bivariate normal distribution. The degree of dependency bivariate data directly affects the shape and curvature structures of the bivariate normal distribution surface. Different parameters, from uncorrelated to highly correlated variables, for the correlation of the bivariate normal distribution based on the data have been examined. The effects of the correlation on the distribution surface been characteristics have analysed individually and collectively. This study presents theoretical results in addition to the results of the simulation and real datasets. The simulation data presents the relationship between the independence of the variables and the uniformity of the κ_{n2} values. The other application is based on the curvature properties of the bivariate normal surface on Covid-19 as real data.

Introduction and literature review

Principal curvatures, principal directions, and normals are used in 3D computer vision to solve basic tasks, such as segmentation, surface classification, surface reconstruction, and registration Krsek et al. (1998). Since the principal curvatures at a surface point are invariant properties, these curvatures are used for surface characterisation. Researchers have studied many methods to estimate the Gaussian and mean curvatures on various surfaces such as faces and human bodies. In Tanaka et al. (1998), the authors presented a correlation-based face recognition approach based on the analysis of principal curvatures. The authors Vemuri et al. (1986) used principal curvatures for shape recognition from range data, while in Deng et al. (2007), the authors used these curvatures for object recognition. In Bedi et al. (1997), two methods for machining complex three-dimensional surfaces on 4 and 5 axis machines are proposed to align the principal axis of the curvature of the machining surface with that of the machined surface to increase the volume of the removed material. Authors Floro-Chason (2001) studied the use of a curvature technique called a multi-beam optical stress sensor. In Sboui et al. (2017), the authors proposed a local surface description around the 3D human body extremities based on the mean of the principal curvature field values on the intrinsic Darcyan parameterisation. Justin et al. (Lev et al. (2017)) presented a new local measure based on the principal curvatures of a point cloud for 3D shape recognition. In Tremblay-Darveau et al. (2018), the authors showed that the principal curvature detection could significantly improve the 3D rendering of relatively noisy ultrasound angiograms without degrading the spatial resolution. Moreover, the authors Wang et al. (2018) proposed a shape- similarity measurement approach based on principal curvatures.

The normal distribution of statistics has great importance in different fields such as economics, medicine, actuarial sciences, and engineering. Azzalini (1985) studied the standard normal-density function and its properties. In Azzalini-Chiogna (2004), the authors gave some exact probability results on the stress-strength model in the case of skew-normal variates. Henze (1986) presented a probabilistic representation of the skew-normal distribution in the normal and truncated normal law. In Branco-Dey (2001), the authors proposed a new class of multivariate skew-elliptical distributions, which bring additional flexibility of modeling skewed data. A class of multivariate unified skew-elliptical distributions is presented in Arellano-Valle-Genton (2010). The authors proposed a new model distribution, known as the bivariate Weibull-Geometric distribution in Kundu-Gupta (2014). In Roozegar-Jafari (2017), the authors expressed the bivariate generalized linear failure rate-power series model as a new class of bivariate distributions. The authors studied a new bivariate distribution and its distributional properties with normal geometric distribution marginal in Mahmoudi-Mahmoodian (2017). In Omair et al. (2018), the authors introduced a new bivariate model and its distributional properties combining negative binomial and geometric distributions. In Jafari et al. (2018), the authors

expressed a new model for bivariate distribution and provided different properties of this distribution by combining the bivariate generalized exponential and power-series distributions. The authors studied a method based on particle swarm optimisation and bivariate normal inverse Gaussian distribution to comprehensively understand the joint period and radius distribution in Kepler exoplanets in Chen–Hung (2019). Additionally, in Jamalizadeh–Balakrishnan (2019), the conditional distribution of a Multivariate Normal Mean-Variance Mixture (MNMVM) distribution is expressed. Shimamoto (2019) examined Japan according to urban park area per capita. The vagueness in the shape-from-shading inference problem is examined in Holtmann-Rice et al. (2018). The author studied the multivariate normal model from the Riemannian geometry perspective by accepting the model as a differentiable manifold in Theil (1984). Colour raster graphics and differential geometry are merged to analyse the surface curvature in Dill (1981).

This study examines the bivariate normal distribution by combining the principal curvatures, Gaussian, and mean curvatures. Additionally, some results based on the relationships between the bivariate normal distribution and surface curvatures are given. Our purpose is to present how to utilise the principal curvatures to examine the characteristic similarities and differences of two normal distributed bivariate real data. This provides a better understanding of real data besides statistical analyses. This study desires to use geometry for statistical inference as a bivariate normal distribution and Covid-19 outbreak.

The remainder of this paper is organised as follows. Preliminary theoretical information about the curvatures of a surface and bivariate normal distribution are briefly reviewed. The surface analysis of the bivariate Gaussian distribution, including the experiments, is presented. Two applications are given to present the relationship between the independence of the variables and the curvature properties of the bivariate normal surface on Covid-19 data. Finally, the paper presents the conclusions.

Theoretical foundations

This section briefly describes the curvatures of a surface and the bivariate normal distribution.

Principal curvatures of a surface

A surface that is a piece of \mathbb{R}^2 in the vicinity of any given point is a subset of \mathbb{R}^3 .

Definition 1: A surface patch $\mathbf{D}: U \to \mathbb{R}^3$ is called regular if it is smooth and the vectors \mathbf{D}_u and \mathbf{D}_v are linearly independent at all points $(u, v) \in U$.

A unit normal of the surface S at a point p can be defined with

$$\mathbf{n}_{\mathbf{D}} = \frac{\mathbf{D}_{\mathbf{u}} \times \mathbf{D}_{v}}{\|\mathbf{D}_{\mathbf{u}} \times \mathbf{D}_{v}\|}$$

Definition 2: Let $\mathbf{D}(u, v)$ be a surface patch of **S** and $\mathbf{\alpha}(t) = \mathbf{D}(u(t), v(t))$ be a unit speed curve in **D**. The normal curvature of the curve is obtained using

 $\kappa_n = L\dot{u}^2 + 2M\dot{u}\dot{v} + N\dot{v}^2,$

in which *L*, *M* and *N* are coefficients of the second fundamental form defined as $L = \mathbf{D}_{uu}\mathbf{n}$, $M = \mathbf{D}_{uv} \cdot \mathbf{n}$, and $N = \mathbf{D}_{vv} \cdot \mathbf{n}$.

Proposition: Let **p** be the point of surface **S**. The principal curvatures, κ_{n1} and κ_{n2} are scalar values, and they are the maximum and minimum of the normal curvature, respectively, at point **p** on surface **S**. The principal curvatures are related to the mean and Gaussian curvatures in a simple manner, with

$$K = \kappa_{n1} \cdot \kappa_{n2} \text{ and } H = \frac{\kappa_{n1} + \kappa_{n2}}{2}.$$

The principal curvatures can be written in terms of the Gaussian and normal curvatures by

$$\begin{cases} \kappa_{n1} = H + \sqrt{H^2 - K} \\ \\ \kappa_{n2} = H - \sqrt{H^2 - K}. \end{cases}$$

The values of the principal and Gaussian curvatures κ_{n1} , κ_{n2} and K play an important role in determining the shape of the surface. At the points of a surface, if $K = \kappa_{n1}$. $\kappa_{n2} > 0$, then these points are elliptical; if $K = \kappa_{n1}$. $\kappa_{n2} < 0$, then these points are hyperbolic, and if $K = \kappa_{n1}$. $\kappa_{n2} = 0$, then these points are parabolic Pressley (2010). For an embedded smooth surface in R^3 , the Gaussian curvature is invariant under the local isometries. The classification of the points on a surface is presented in Table 1.

Table 1

Classification of points on a surface according to principal curvatures κ_{n1} and κ_{n2}

| | $\kappa_{n1} < 0$ | $\kappa_{n1} = 0$ | $\kappa_{n1} > 0$ |
|-------------------|---------------------|-------------------|---------------------|
| $\kappa_{n2} < 0$ | Concave ellipsoid | Concave cylinder | Hyperboloid surface |
| $\kappa_{n2} = 0$ | Concave cylinder | Plane | Convex cylinder |
| $\kappa_{n2} > 0$ | Hyperboloid surface | Convex cylinder | Convex ellipsoid |

Bivariate normal distribution

A vector-valued random variable $X = [X_1, ..., X_p]^T$ is said to have a multivariate normal distribution with a mean vector $\overline{\mu}$ and a covariance matrix Σ if its probability density function is given by

$$\begin{cases} f(\mathbf{x}; \boldsymbol{\mu}, \boldsymbol{\Sigma}) = \\ \frac{1}{(2\pi)^{\frac{d}{2}} |\boldsymbol{\Sigma}|^{\frac{1}{2}}} \exp\left\{-\frac{1}{2}(\mathbf{x}-\boldsymbol{\mu})^T \boldsymbol{\Sigma}^{-1}(\mathbf{x}-\boldsymbol{\mu})\right\}. \end{cases}$$

In the case of multivariate normal density, the argument of the exponential function, $\left(-\frac{1}{2}(\mathbf{x}-\boldsymbol{\mu})^T \Sigma^{-1}(\mathbf{x}-\boldsymbol{\mu})\right)$ is a quadratic form of the vector variable **x**. It is

known that the covariance matrix Σ is positive definite, for any vector $\mathbf{x} \neq \boldsymbol{\mu}$. If $(\mathbf{x} - \boldsymbol{\mu})^T \Sigma^{-1} (\mathbf{x} - \boldsymbol{\mu}) > 0$ and $\left(-\frac{1}{2} (\mathbf{x} - \boldsymbol{\mu})^T \Sigma^{-1} (\mathbf{x} - \boldsymbol{\mu})\right) < 0$, then the argument of the exponential function is a downward-opening quadratic bowl. In the case of p = 2, the bivariate normal density has the form:

$$f(\mathbf{x}; \boldsymbol{\mu}, \boldsymbol{\Sigma}) = \frac{1}{(2\pi\boldsymbol{\sigma}_1\boldsymbol{\sigma}_2)(1-\rho)^{\frac{1}{2}}} \exp\left(-\frac{z}{2(1-\rho)^2}\right)$$

where σ is standard deviation, and

$$\mathbf{z} \equiv \frac{(\mathbf{x}_1 - \boldsymbol{\mu}_1)^2}{\sigma_1^2} - \frac{2\rho(\mathbf{x}_1 - \boldsymbol{\mu}_1)(\mathbf{x}_2 - \boldsymbol{\mu}_2)}{\sigma_1 \sigma_2} + \frac{(\mathbf{x}_2 - \boldsymbol{\mu}_2)}{\sigma_2^2}$$

and also

$$\rho \equiv cor(\mathbf{x}_1, \mathbf{x}_2) = \frac{\sigma_{1,2}}{\sigma_1 \sigma_2}$$

which is the correlation of x_1 and x_2 .

As known, the contours of a bivariate normal distribution are elliptical. Figure A1 in the Appendix illustrates three cases: positive, negative, and zero correlations.

Surface curvature analysis of bivariate Gaussian distribution

The surface analysis of the bivariate Gaussian distribution is examined using the maximum, minimum, Gauss, and mean curvatures. In this study, 10,000 data are generated based on the mean vector and covariance matrix as

$$\mathbf{\mu} = \begin{pmatrix} 0 \\ 0 \end{pmatrix} \text{ and } \Sigma = \begin{pmatrix} 1 & \rho \\ \rho & 1 \end{pmatrix}.$$

Then, the maximum, minimum, Gaussian, and mean curvatures of these data are computed according to the correlation values of 0.0; 0.1; 0.3; 0.5; 0.7; 0.75; 0.80; 0.85; 0.90, and 0.95. Meanwhile, since the same results are obtained for the negative correlation values, only positive correlation values are considered in this study. The classification of points of the Bell-shaped surfaces, which belong to the data produced using different correlation values, is presented in Table 2. Since different values of κ_{n1} corresponding to the values of $\kappa_{n2} = 0$ and $\kappa_{n2} > 0$ are equal to zero, we did not show these values in Table 2.

For all the examined values of ρ , κ_{n2} is obtained as negative, whereas κ_{n1} can receive positive and negative values. At all points on the Bell-shaped surface, the change rate of κ_{n1} sign is 40% and 60%, which is not affected by the change in ρ . Next, the relationship between the Gaussian curvature and correlation, as well as the mean curvature and correlation, can be seen in Figures A2 and A3 in the Appendix. The locations where κ_{n1} is calculated as positive and negative compared to the regions of the Bell-shaped surface. Additionally, at the centre of and around the distribution, while *K* and *H* take positive values, they turn into negative values in the skirts and exhibit a proportionally unchanged appearance with ρ . Figures A4 and A5 in the

Appendix show the distribution of κ_{n1} and κ_{n2} , and in the weak correlations, the distribution of κ_{n1} is skewed to the left, while the right skewness changes gradually for strong correlation values. Conversely, the distribution of κ_{n2} tends to be uniform for small correlation values, while it becomes skewed left for high correlation values. The marginal plots of κ_{n1} and κ_{n2} are shown in Figure A6 in the Appendix. The variation between κ_{n1} and κ_{n2} depending on the different values of ρ is shown in Figure A6.

Table 2

| ρ | $\kappa_{n1} < 0 \& \kappa_{n2} < 0 (\%)$ | $\kappa_{n1} > 0 \& \kappa_{n2} < 0 (\%)$ | Total |
|-------|---|---|--------|
| 0.00 | 39.57 | 60.43 | 100.00 |
| 0.10 | 39.90 | 60.10 | 100.00 |
| 0.30 | 39.65 | 60.35 | 100.00 |
| 0.50 | 40.35 | 59.65 | 100.00 |
| 0.70 | 39.52 | 60.48 | 100.00 |
| 0.75 | 39.35 | 60.65 | 100.00 |
| 0.80 | 39.76 | 60.24 | 100.00 |
| 0.85 | 40.21 | 59.79 | 100.00 |
| 0.90 | 39.26 | 60.74 | 100.00 |
| 0.95 | 39.14 | 60.86 | 100.00 |
| Total | 39.67 | 60.33 | 100.00 |

Classification of points of the Bell-shaped surfaces according to different correlation values

Figure 1

The change between the principal curvatures κ_{n1} and κ_{n2} for uncorrelated variables



Moreover, in the case of independent variables, the principal curvatures κ_{n1} and κ_{n2} change cubically. The obtained cubic regression polynomial is shown in Figure 1.

As ρ values become greater than zero, the interval changes for both the principal curvatures κ_{n1} and κ_{n2} increase (see Figure 2). Additionally, the average curvature κ_{n2} decreases, whereas the average curvature κ_{n1} increases. Figure 2 shows that the change in κ_{n2} is realised over a wider range. However, it is noteworthy that these changes are not linear.





Figure 2 also shows that the principal curvatures κ_{n1} and κ_{n2} change in opposite directions and nonlinearly as the correlation increases. Additionally, it is noteworthy that the principal curvature κ_{n2} changes over a wider range.

Application

In this section we stated that the distribution of κ_{n1} is skewed to the left for small correlations. Simultaneously, it tends to the right for strong correlations, and κ_{n2} tends to be distributed uniformly for small correlations but skewed left for high correlations. Meanwhile, as seen from the graph in Figure A5 in the Appendix, the uniform case for κ_{n2} corresponds to the independence of the normal variables. To show the relationship between the independence of the variables and the uniformity of the κ_{n2} values, we designed two examples.

Example 1

Random samples are generated from the bivariate normal distribution of sizes n = 50, n = 250, and n = 1,000 and different correlation values. Pearson's correlation coefficients and 95% confidence intervals are calculated for each sample. Additionally, κ_{n2} values are obtained using the same data. To investigate the distribution of κ_{n2} values, we performed a goodness-of-fit test using Kolmogorov-Smirnov test statistics. The critical region is based on the distribution of the test statistics under the uniformity assumption is true. The results are summarised in Table 3.

Table 3

| Value Correlation Kolmogroup-Smirnov test r 95%CI Value Result n=50 n=50 ro=0,0 -0.158 -0.418 0.126 0.071 Accept p>0.05 ro=0,3 0.146 -0.322 0.234 0.06774 Accept p>0.05 ro=0,5 0.569 0.345 0.731 0.09678 Accept p>0.05 ro=0,7 0.821 0.703 0.895 0.15594 Accept p>0.05 ro=0,80 0.824 0.708 0.897 0.13529 Accept p>0.05 ro=0,85 0.846 0.743 0.911 0.18379 Accept p>0.05 ro=0,90 0.897 0.825 0.941 0.24034 Reject p<0.05 ro=0,0 -0.023 -0.147 0.101 0.03202 Accept p>0.05 ro=0,1 0.074 -0.051 0.196 0.05561 Accept p>0.05 ro=0,5 0.536 0.441 < | | | | | | | |
|---|---------|-------------|--------|-------------------------|---------|--------|--------|
| r95%CIValueResultn=50ro=0,0-0.158-0.4180.1260.071Accept $p>0.05$ ro=0,1-0.048-0.3220.2340.06774Accept $p>0.05$ ro=0,30.146-0.1380.4080.15346Accept $p>0.05$ ro=0,70.8210.7030.8950.15594Accept $p>0.05$ ro=0,70.8210.7030.8950.15594Accept $p>0.05$ ro=0,800.8240.7080.8970.13529Accept $p>0.05$ ro=0,850.8460.7430.910.18379Accept $p>0.05$ ro=0,900.8970.8250.9410.24034Reject $p<0.05$ ro=0,0-0.023-0.1470.1010.03202Accept $p>0.05$ ro=0,0-0.024-0.0440.1960.05561Accept $p>0.05$ ro=0,0-0.023-0.1470.1010.03202Accept $p>0.05$ ro=0,0-0.024-0.0440.6190.10353Reject $p<0.05$ ro=0,0-0.0230.0770.3150.0594Accept $p>0.05$ ro=0,50.5360.4410.6190.10353Reject $p<0.05$ ro=0,50.5360.4410.6190.10353Reject $p<0.05$ ro=0,70.6 | Value | Correlation | | Kolmogorov-Smirnov test | | | |
| $\begin{array}{c c c c c c c c c c c c c c c c c c c $ | value | r | 95% | юCI | Value | Re | sult |
| $\begin{array}{c c c c c c c c c c c c c c c c c c c $ | | n=50 | | | | | |
| $\begin{array}{c c c c c c c c c c c c c c c c c c c $ | ro=0,0 | -0.158 | -0.418 | 0.126 | 0.071 | Accept | p>0.05 |
| $\begin{array}{c c c c c c c c c c c c c c c c c c c $ | ro=0,1 | -0.048 | -0.322 | 0.234 | 0.06774 | Accept | p>0.05 |
| $\begin{array}{c c c c c c c c c c c c c c c c c c c $ | ro=0,3 | 0.146 | -0.138 | 0.408 | 0.15346 | Accept | p>0.05 |
| $\begin{array}{c c c c c c c c c c c c c c c c c c c $ | ro=0,5 | 0.569 | 0.345 | 0.731 | 0.09678 | Accept | p>0.05 |
| $\begin{array}{c c c c c c c c c c c c c c c c c c c $ | ro=0,7 | 0.821 | 0.703 | 0.895 | 0.15594 | Accept | p>0.05 |
| $\begin{array}{c c c c c c c c c c c c c c c c c c c $ | ro=0,75 | 0.742 | 0.584 | 0.846 | 0.15951 | Accept | p>0.05 |
| $\begin{array}{c c c c c c c c c c c c c c c c c c c $ | ro=0,80 | 0.824 | 0.708 | 0.897 | 0.13529 | Accept | p>0.05 |
| $\begin{array}{c c c c c c c c c c c c c c c c c c c $ | ro=0,85 | 0.846 | 0.743 | 0.91 | 0.18379 | Accept | p>0.05 |
| $\begin{array}{c c c c c c c c c c c c c c c c c c c $ | ro=0,90 | 0.897 | 0.825 | 0.941 | 0.24034 | Reject | p<0.05 |
| $\begin{array}{c c c c c c c c c c c c c c c c c c c $ | ro=0,95 | 0.959 | 0.928 | 0.977 | 0.28453 | Reject | p<0.05 |
| $\begin{array}{c c c c c c c c c c c c c c c c c c c $ | | n=250 | | | | | |
| $\begin{array}{c c c c c c c c c c c c c c c c c c c $ | ro=0,0 | -0.023 | -0.147 | 0.101 | 0.03202 | Accept | p>0,05 |
| $\begin{array}{c c c c c c c c c c c c c c c c c c c $ | ro=0,1 | 0.074 | -0.051 | 0.196 | 0.05561 | Accept | p>0,05 |
| $\begin{array}{c c c c c c c c c c c c c c c c c c c $ | ro=0,3 | 0.199 | 0.077 | 0.315 | 0.0594 | Accept | p>0,05 |
| $\begin{array}{c c c c c c c c c c c c c c c c c c c $ | ro=0,5 | 0.536 | 0.441 | 0.619 | 0.10353 | Reject | p<0.05 |
| $\begin{array}{c c c c c c c c c c c c c c c c c c c $ | ro=0,7 | 0.668 | 0.593 | 0.732 | 0.15862 | Reject | p<0.05 |
| $\begin{array}{c c c c c c c c c c c c c c c c c c c $ | ro=0,75 | 0.728 | 0.664 | 0.782 | 0.17356 | Reject | p<0.05 |
| $\begin{array}{c c c c c c c c c c c c c c c c c c c $ | ro=0,80 | 0.762 | 0.704 | 0.809 | 0.18518 | Reject | p<0.05 |
| $\begin{array}{c c c c c c c c c c c c c c c c c c c $ | ro=0,85 | 0.848 | 0.809 | 0.88 | 0.1922 | Reject | p<0.05 |
| $\begin{array}{c c c c c c c c c c c c c c c c c c c $ | ro=0,90 | 0.924 | 0.903 | 0.94 | 0.22662 | Reject | p<0.05 |
| $\begin{array}{c c c c c c c c c c c c c c c c c c c $ | ro=0,95 | 0.942 | 0.926 | 0.954 | 0.2668 | Reject | p<0.05 |
| $\begin{array}{c c c c c c c c c c c c c c c c c c c $ | | n=1,000 | | | | | |
| $ \begin{array}{c c c c c c c c c c c c c c c c c c c $ | ro=0,0 | -0.034 | -0.096 | 0.0228 | 0.022 | Accept | p>0,05 |
| $\begin{array}{c c c c c c c c c c c c c c c c c c c $ | ro=0,1 | 0.077 | 0.016 | 0.139 | 0.04903 | Reject | p<0.05 |
| $\begin{array}{c c c c c c c c c c c c c c c c c c c $ | ro=0,3 | 0.255 | 0.196 | 0.312 | 0.07695 | Reject | p<0.05 |
| $\begin{array}{c c c c c c c c c c c c c c c c c c c $ | ro=0,5 | 0.467 | 0.417 | 0.514 | 0.11577 | Reject | p<0.05 |
| $ \begin{array}{c c c c c c c c c c c c c c c c c c c $ | ro=0,7 | 0.684 | 0.65 | 0.716 | 0.17822 | Reject | p<0.05 |
| $ \begin{array}{c c c c c c c c c c c c c c c c c c c $ | ro=0,75 | 0.731 | 0.7 | 0.758 | 0.17669 | Reject | p<0.05 |
| $ \begin{array}{c c c c c c c c c c c c c c c c c c c $ | ro=0,80 | 0.81 | 0.787 | 0.83 | 0.18844 | Reject | p<0.05 |
| ro=0,90 ro=0,95 0.954 0.948 0.912 0.21422 Reject p<0.05 0.959 0.2545 Reject p<0.05 | ro=0,85 | 0.848 | 0.83 | 0.865 | 0.20358 | Reject | p<0.05 |
| ro=0,95 0.954 0.948 0.959 0.2545 Reject p<0.05 | ro=0,90 | 0.9 | 0.888 | 0.912 | 0.21422 | Reject | p<0.05 |
| | ro=0,95 | 0.954 | 0.948 | 0.959 | 0.2545 | Reject | p<0.05 |

Pearson's correlation coefficients, 95% confidence intervals and Kolmogorov-Smirnov test results of κ_{n2} values

It is seen from Table 3 that for n = 50, the independence of the variables is not exactly supported by the uniformity test but for n = 250, and n = 1,000 independency of the variables support gradually by the uniformity test. For n = 1,000 and $\rho = 0.0$, the correlation test states that variables are independent and Kolmogorov-Smirnov the test states that κ_{n2} is distributed uniformly. On the other hand, for n = 1,000 where ρ values are greater than 0.0, the dependency characteristic of the original variables produces non-uniform κ_{n2} values.

Example 2

The curvature properties of a bivariate normal surface on real data will be explained in this section. We have been struggling with Covid-19 for more than two years, and there are different studies on this topic. Severe acute respiratory syndrome (SARS) and Covid-19 are compared in terms of similarities and differences in Wilder-Smith et al. (2020). Foldvary (2021) developed a MATLAB tool to analyse and monitor the spread of COVID-19. The author examined the spread based on a geostatistical analysis using MATLAB. A project focused on the spatial correlations among socialdemographic, social-demographic, weather information, and atmospheric pollution data with lethality rates and contagion of Covid-19 in Ireland Aghdam-Crowley (2020). The pattern of COVID-19 victims is examined in terms of gender, age, and medical problems. The authors used K-Means and two-step clustering methods for the analysis in Gholipour et al. (2021). The authors studied the effect of Covid-19 on small and medium enterprises in Hungarian Nyikos et al. (2021). In this study, we collect 60-day data on deaths, intubated patients, and daily confirmed virus case numbers given sequentially and regularly between 27 March 2020 and 25 May 2020 from the Covid-19 data of the Turkey Ministry of Health. Many observations of biological processes and characteristics tend to follow a normal distribution. However, the individual case distribution characteristics obtained for this period were observed to have different characteristics than the normal distribution, but using these data: two ratios assumed to be bivariate normally distributed were calculated. The first ratio is the daily death number/case number, and the other is the percentage of the daily intubated patients number/case number. The change of these two ratios is given in Figure 3.

As shown in Figure 3, the Pearson correlation value between the two ratios was obtained as 0.681, and it was observed that there was a moderate, strong linear relationship between them. For the marginal distributions of the two rates considered, the normality test was applied using probability-plot and Kolmogorov-Smirnov test statistics, and the results are shown in Figure A7 in the Appendix.



Table 4



The histograms of both death to case and intubated patients to a) case ratios, b) the normality test for death to case ratio and c) the normality test for intubated patients to case ratio.

Figure A7 in the Appendix shows that the normality assumption is valid for both ratios (p > 0.15). The average death to case percentage for the 60-day period was 2.874, and the average for intubated patients to case percentage was 31.57. The covariance structures for the mentioned rates are presented in Table 4.

| Covariances of ratios | | | | |
|----------------------------------|---------------------|----------------------------------|--|--|
| | Death to case ratio | Intubated patients to case ratio | | |
| Death to case ratio | 0.612343 | - | | |
| Intubated patients to case ratio | 4.259030 | 63.923556 | | |

Within the scope of this study, κ_{n1} and κ_{n2} curvatures are calculated for the Covid-19 mentioned ratios, and are shown in Figure A8 in the Appendix.

The histogram of Gauss and mean curvatures together with κ_{n1} and κ_{n2} are shown in Figure A9 in the Appendix.

As there is a correlation between the examined rates, it is seen from Figure A9 in the Appendix that the uniform distribution structure is disrupted and the right-skewed distribution structure appears for κ_{n1} and the opposite direction skewed distribution character appears for κ_{n2} . In the context of this study, Figure A6 (e) option, which is obtained based on the simulation results for $\rho = 0.70$, is the closest

to the correlation value 0.681 which is calculated for the rates. The marginal distributions for κ_{n1} and κ_{n2} calculated according to the Covid-19 ratios that are obtained in this study are given in Figure A10 in the Appendix.

The resulting structure for the marginal curvature distributions and the simulation case for $\rho = 0.7$ are given in Figure A10.

Conclusion

Principal curvatures, Gaussian and mean curvatures at a point on a surface are used for surface characterisation. Statistics is crucial for making some discoveries and predictions using data in many fields. This study presented a novel approach to the bivariate normal distribution and expressed the effectiveness of the correlation of the bivariate normal distribution related to principal curvatures, Gaussian, and mean curvatures of the Bell-shaped surface.

The principal curvatures in the bivariate normal distribution were examined for ten different correlation levels ranging from non-correlated to high correlation. The principal curvature κ_{n1} shows a uniform distribution in the absence of correlation. However, as the correlation increases, it has a left-skewed distribution characteristic. The principal curvature κ_{n2} , unlike the first one, changes from a right-skewed distribution characteristic in the absence of correlation to a left-skewed distribution characteristic as in the first curve when the correlation increases. The change between the principal curvatures is observed in a cubic polynomial model in the absence of correlation. Additionally, the proportional change of the principal curvature κ_{n2} producing negative values for the positive and negative values of the principal curvature κ_{n1} were independent of the correlation characteristics of the bivariate data. The tests show that the bivariate normal distribution surface varies between the concave ellipsoid and hyperboloid structures. The distribution has systematic characteristics arising from the correlation structure at the skirt and top parts of the distribution. Finally, we provided two applications to show the relationship between the independence of the variables and the uniformity of the κ_{n2} values and to present the curvature properties of the bivariate normal surface on the Covid-19 real data. Similarities and differences in the Covid-19 process can be realised between two places according to the principal curvatures.

This study reveals the change between curvature characteristics and variable dependence in bivariate distribution and provides a new perspective on recognising distribution. Hence, the characteristic similarities and differences between two normally distributed bivariate real data can be examined using only the principal curvatures.

Appendix

Figure A1



Figure A2





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(Continued.)





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Figure A3



0.2

0.15

0.1

0.05

0





d) ho=0.5

-5 -4

-2

b) $\rho = 0.1$



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(Continued.)

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Figure A4



The distribution of the values of κ_{n1} for the values of ρ

Figure A5

The distribution of the values of κ_{n2} for the values of ρ



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e)









(The figure continued next page.)





Figure A7

The histograms and normality test for both death to case and intubated patients to case ratios

a) Case ratios

b) The normality test for death to case ratio



c) The normality test for intubated patients to case ratio





Figure A8

κ_{n1} and κ_{n2} curvatures of the ratios that are death to case and intubated patients to case



Figure A9

The histogram of Gauss and mean curvatures together with the principal curvatures κ_{n1} and κ_{n2}



Figure A10

The histogram of Gauss and mean curvatures together with the principal curvatures κ_{n1} and κ_{n2} The marginal distributions of b) The relationship between $\kappa n1$ a) κ_{n1} and κ_{n2} and κ_{n2} for $\rho = 0.7$ ρ=0.7 Marginal Plot of Kn1 vs Kn2 of Covid-19 Data 0,3 : 0,2 5 0,1 ; 0,0 2 2 -0,1 -0,4 Kn2

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